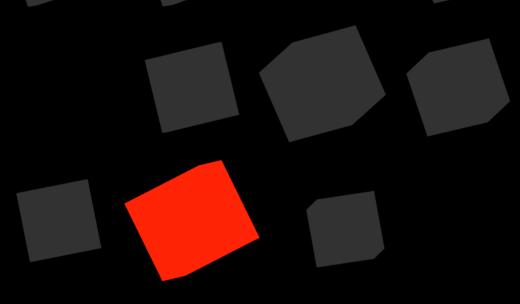


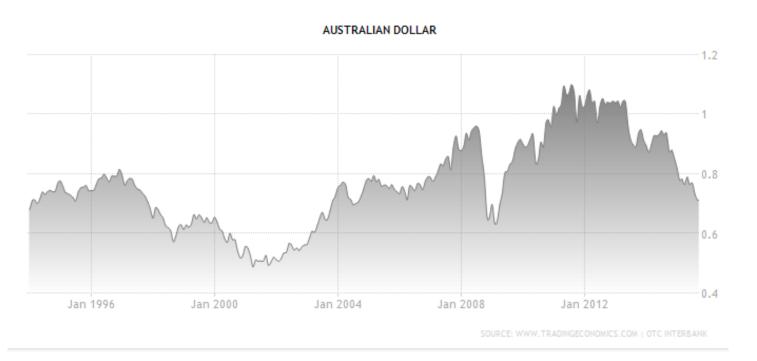
37161 Probability and Random Variables

Lecture 9



UTS CRICOS 00099F

 A time series is a collection of datapoints recording observations made at successive intervals of some time period.



• For example, the graph here show the conversion of the \$US to \$AUS over the past 20 years.



ÖUTS

- Analysing such time series can be exceptionally complex:
- Does the distribution of possible future datapoints depend only on the current level?
- Is there "momentum" in the system i.e. should past performance be used as an predictor as well?
- Can other factors outside the series of observations also be useful predictors?



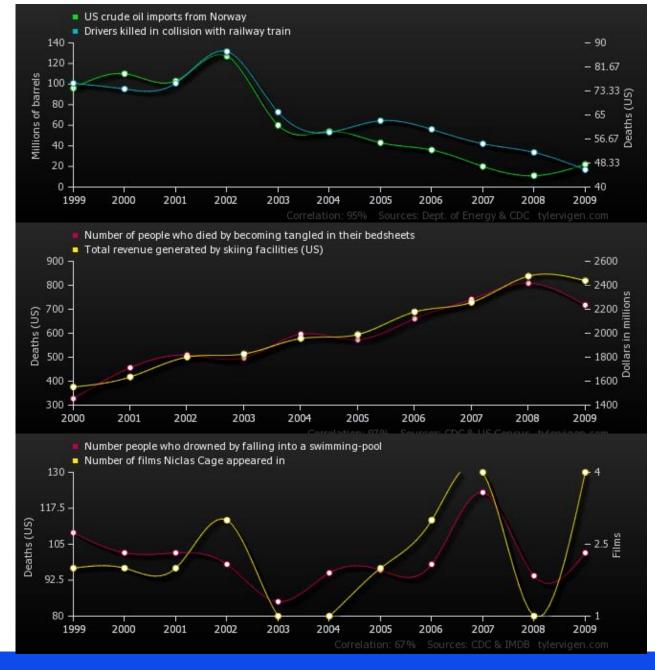
 Most time series analyses are concerned with making predictions about possible future observations, based on both the past trends and any other relevant factors.



- Typical analyses include weather forecasting, economic/financial modelling, climate modelling etc.
- Understanding what is and is not relevant from observed datasets in predicting future trends is often highly debated.



 The website Spurious Correlations (<u>http://www.tylervigen.com/spurious-</u> <u>correlations</u>) lists a few humourous examples of cases when, despite strong correlation, a past trend might not be such a good predictor of future results...



Markov Chains

ÖUTS

- Consider a sequence of random variables $\{X_{0,}X_{1,}X_{2,}...\}$ each of which takes one of a finite or countable number of values.
- We call the set of possible values for each variable the states of the system, *S*
- For many problems of interest, we have the property that

$$P(X_{n+1} = j | X_n = i, X_{n-1} = i_{n-1}, X_{n-2} = i_{n-2}, \dots, X_0 = i_0) = P(X_{n+1} = j | X_n = i)$$

for all possible *n* and all possible *j*, *i*, *i*_{*n*-1}, *i*_{*n*-2}, ..., *i*₀ \in S

• In other words, the only necessary predictor of future states is the current one.

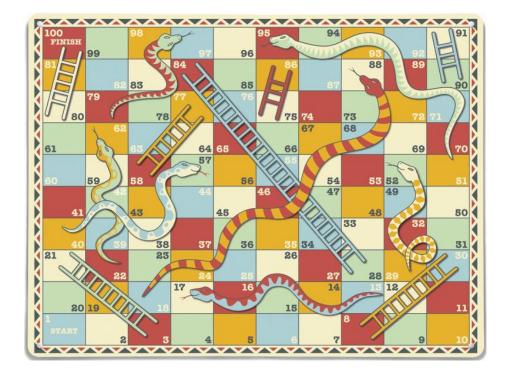


Andrey Markov (1856 - 1922)

• Such a system is called a Markov Chain.

Markov Chains

• For example, with a standard game of snakes and ladders, with squares numbered 1-100, the square that a given player is on at the end of each turn forms a Markov Chain with states $S = \{1, 2, 3, ..., 100\}$



• This is because if we were told the full history of all a player's positions, for example

$$X_{1} = 41, X_{0} = 36, X_{3} = 31, X_{3} = 26, X_{3} = 19, X_{3} = 15, X_{4} = 12, X_{3} = 7, X_{2} = 4, X_{1} = 2, X_{3} = 10, X_{3} = 10, X_{4} = 12, X_{4} =$$

and asked to calculate the distribution of X_{12} , the only relevant information is that $X_{11} = 41$.

• The possible moves from square 41 do not depend on how the player got there.

Markov Chains

- Which of the following might reasonably be modelled with Markov Chains through time?:
- The proportion of people in a country's population who have green eyes .
- Markov Chain. No reason to assume proportion is more likely to increase or decrease in time.
- The proportion of people in a country's population who have a harmful genetic condition.
- Not Markov Chain. If the condition is harmful, fewer people with it may be able to reproduce.
- The volume of water in a bucket, given that 1 Litre is added per minute with none removed.
- Markov Chain. We could know future volumes from knowing only current volume.
- The amount of money a gambler has playing a sequence of independent identical games.
- Depends... If it's a game of pure chance, then would be a Markov Chain. If there is skill involved, then prior states would give evidence of the player's possible future performance.

Random Walks

- One of the most commonly studied types of Markov Chain is a **random walk**.
- A random walk is a sequence defined on the integers such that a probability P_{ij} is assigned to the probability of the (n+1)th state being j, given that the nth state is i.



- Consider for example the simple game of repeatedly betting \$1 on the flip of a fair coin to win \$2 if the result is Heads and \$0 if the result is Tails.
- This defines a Markov Chain since, for example, the chance of going from \$15 to \$16 is completely independent of any previous flips, i.e. whether you started with \$1 and kept winning until you got \$15 or whether you started with \$100 and kept losing until you got \$15.

A Simple Random Walk

- For example, consider a game whereby a player bets \$1 on the flip of a fair coin. He/she wins \$2 is the result is Heads and \$0 if the result is Tails.
- Staring with \$10, he/she bets repeatedly until he/she either has \$0 or \$20.



- We can set up a **difference equation** in W_k for the probability that he eventually leaves in profit, given that he has $k \in \{0, 1, 2, \dots, 20\}$ dollars at a given time.
- The probability can be broken down conditional on the next game. Given that he/she has \$k, he/she has a 50% chance of having (k+1) dollars after the next game and a 50% chance of having \$(k 1).

A Simple Random Walk

- We know that if we have two events B_1 and B_2 which together form a partition of the sample space, i.e. $P(B_1 \cup B_2) = 1$ and $P(B_1 \cap B_2) = 0$ then, for any event A, $P(A) = P(A \cap B_1) + P(A \cap B_2)$
- Applying this same idea for each individual game, we have that the probability that a player wins overall is equal to the probability that he/she wins the next game and wins overall plus the probability that he/she loses the next game but wins overall.
- This is $W_k = P(\text{loses next game}) \times W_{k-1} + P(\text{wins next game}) \times W_{k+1}$
- For a game where the probability of winning and losing \$1 are each 50%, the difference equation is $W_k = \frac{1}{2}W_{k+1} + \frac{1}{2}W_{k-1}$
- This is solved, subject to boundary conditions $W_0 = 0$ since if he ever has \$0, he cannot win overall (probability 0) and $W_{20} = 1$ since if he reaches \$20

• A first order difference equation (i.e. one in which the highest and lowest numbered terms differ by exactly 1) with constant coefficients can be easily solved by inspection.

• For example, if $C_{n+1} = 2C_n$, we can easily see that $C_{n+2} = 2(C_{n+1}) = 2^2C_n$ and $C_{n+3} = 2(C_{n+2}) = 2^3C_n$.

• In general, $C_{n+1} = 2C_n$ is solved by $C_n = K2^n$ for some constant K.

• Similarly, for constant *M*, $C_{n+1} = MC_n$ is solved by $C_n = KM^n$.

- How in general do we solve a **second order difference equation** (i.e. one in which the highest and lowest numbered terms differ by exactly 2)?
- To solve, for example $C_{n+2} 7C_{n+1} + 10C_n = 0$, we look for a solution of the form $C_n = KM^n$ for constants *K* and *M*.
- If we get two different values for *M*, then the general solution is $C_n = K_1 M_1^n + K_2 M_2^n$.
- If we get one repeated value for *M*, then the general solution is $C_n = K_1 M^n + K_2 n M^n$.
- Compare this with solving second order differential equations of the form $\frac{d^2y}{dx^2} + A\frac{dy}{dx} + y = 0$ to find solutions of the form $y = K_1 e^{m_1 x} + K_2 e^{m_2 x}$ (or $y = K_1 e^{mx} + K_2 x e^{mx}$ if one repeated value of *m* is found.)

- To solve, for example $C_{n+2} 7C_{n+1} + 10C_n = 0$, we look for a solution of the form $C_n = KM^n$ for constants *K* and *M*.
- If such a solution exists, $C_{n+1} = KM^{n+1} = (KM^n)M = C_nM$ and $C_{n+2} = KM^{n+2} = (KM^n)M^2 = C_nM^2$.
- Substituting these in gives $M^2C_n 7MC_n + 10C_n = 0$ or $C_n[M^2 7M + 10] = 0$.
- Apart from the trivial solution $C_n = 0$, we have that $M^2 7M + 10 = (M 5)(M 2) = 0$ so M = 5 or 2.
- The general solution is therefore $C_n = K_1 5^n + K_2 2^n$.
- Boundary conditions are needed to work out the two constants.

∛UTS

- Returning to our initial problem of $W_k = \frac{1}{2}W_{k+1} + \frac{1}{2}W_{k-1}$, we seek solutions of the form $W_k = AM^k$.
- We find *M* by solving $W_k = \frac{1}{2}MW_k + \frac{1}{2}\frac{W_k}{M}$ or $M^2 2M + 1 = 0$.
- As this gives (M-1)(M-1) = 0, we solve to find that M = 1.
- This gives a repeated value for *M*, so $W_k = A_1 1^k + A_2 k 1^k$ or $W_k = A_1 + A_2 k$.
- The boundary conditions of $W_0 = 0$ and $W_{20} = 1$ give that $W_0 = A_1 = 0$ and $W_{20} = A_1 + 20A_2 = 1$.
- Together, we get that $W_k = \frac{k}{20}$ so, for example, the probability of winning \$20 rather than \$0 after starting with \$10 is, perhaps obviously 50%. This is because the starting position is equally spaced from both boundaries.

Asymmetric Random Walks

- The cases when the probabilities of winning and losing each game are not equal is a little simpler to solve as it does not involve a repeated value of *M*.
- For example, consider a European style roulette wheel. For a \$1 bet, a gambler gets \$2 back with probability $\frac{18}{37}$ and \$0 back with probability $\frac{19}{37}$.
- If a gambler plays repeatedly until first getting to \$*D* or \$0, then the difference equation which describes his overall chance of winning, given that he has \$*k* at a given point is

$$W_k = \frac{18}{37}W_{k+1} + \frac{19}{37}W_{k-1}$$
 with boundary conditions $W_0 = 0$ and $W_D = 1$



Asymmetric Random Walks

- We find M via $18M^2 37M + 19 = 0$ so (18M 19)(M 1) = 0.
- $M = \frac{19}{18}$ or 1 hence the general solution is $W_k = A_1 \left(\frac{19}{18}\right)^k + A_2 \left(1\right)^k = A_2 + A_1 \left(\frac{19}{18}\right)^k$
- The boundary conditions give

$$W_0 = A_2 + A_1 = 0$$
 and $W_D = A_2 + A_1 \left(\frac{19}{18}\right)^D$





Asymmetric Random Walks

- The solution of $W_k = \frac{18}{37} W_{k+1} + \frac{19}{37} W_{k-1}$ with boundary conditions $W_0 = 0$ and $W_D = 1$ is therefore $W_k = \frac{1 - \left(\frac{19}{18}\right)^k}{1 - \left(\frac{19}{18}\right)^D}$. • So, even though the odds are very nearly fair ($\frac{18}{37} \approx 48.65\%$), if starting with \$100 and playing until reaching either \$200 or \$0, the probability of ending in profit is only

$$W_{100} = \frac{1 - \left(\frac{19}{18}\right)^{100}}{1 - \left(\frac{19}{18}\right)^{200}} \approx 0.45\%$$

For a truly fair game, this would be 50%, of course. ٠

Simple Random Walks: How Many Games?

- We have calculated the probabilities of either quitting in profit or bankrupt, but how many games should the player expect to have to play before concluding his bets, assuming a fair game?
- Clearly, if he starts at \$10 and leaves at \$0 or \$20, he must play a minimum of 10 games (either 10 losses or 10 wins) but could keep playing a lot longer than that.
- Let C_k be a "counter" of how many games he expects to play until he finishes, given that he has \$k.
- Again, conditioning on the next move, he either wins \$1 (with probability 0.5) or loses \$1 (with probability 0.5). Either way, the "counter" moves on by one.
- We therefore get $C_k = \frac{1}{2} [C_{k+1} + 1] + \frac{1}{2} [C_{k-1} + 1]$.

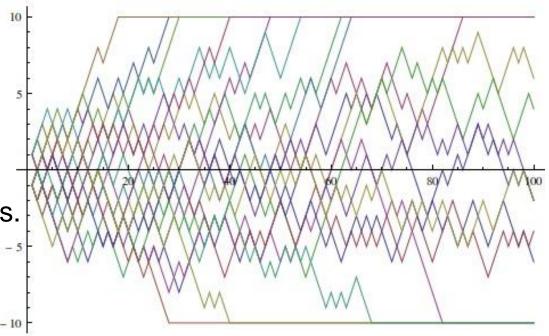
Simple Random Walks: How Many Games?

•
$$C_{k} = \frac{1}{2} [C_{k+1} + 1] + \frac{1}{2} [C_{k-1} + 1]$$
.

- The boundary conditions are $C_0 = 0, C_{20} = 0$ since he requires no games to quit if he starts with \$0 or \$20.
- The solution of this is found via a nonhomogeneous difference equation with repeated roots beyond the scope of this subject. (For completeness, the solution is $C_k = 20k k^2$)
- Again, perhaps unsurprisingly, this confirms that the expected number of games is maximised when k = 10 i.e. when the starting position is furthest from either boundary.
- What do we do in such cases when we know the mechanisms of a model, but are unable to calculate an analytical solution?

Simulation

- Computer packages can be used to generate large numbers of realisations of random systems.
- Expectations of random variables can then be estimated from averaging the outcomes of these simulations.



- For example, the following is the output of 20 realisations of a random walk beginning with \$10 and playing until first reaching \$20 (\$10 profit) or \$0 (shown here as \$ –10 profit.)
- As well as being able to estimate probabilities of winning etc. expected numbers of games or other more obscure properties can be estimated.