Regression and Linear Models (37252) Lecture 5 - Multiple Linear Regression II

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These notes are based, in part, on earlier versions prepared by Dr Ed Lidums and Prof. James Brown.

Topics:

- model selection
 - problem setup
- school performance example
 - setup
 - selection from all possible models
 - forward selection
 - backward elimination
 - stepwise regression

5. Model selection

We have now been introduced to the basics of linear regression and know enough to

- 1 propose a model
- 2 build a model
- 3 analyse in term of fit and satisfaction of assumptions.

Of these we have had a good look at the second and third on the list.

Our exposure to the first has been limited to examples in multiple regression requiring the selection of independent variables from a limited number of alternatives, and to simple linear regression where the choice was even more straightforward.

In this lecture we look more closely at this and develop tools for model selection in more complicated settings.

For more details see Chapter 15 of Draper and Smith (1998).

In practical situations, the models we design have to satisfy competing principles.

On the one hand, we want them to be as accurate as possible. On the other, we would like them to be a simple as possible.

The need for accuracy is obvious. The need for simplicity, essentially, relates to cost, be it the cost of equipment, staff or even delay.

Even if simplicity was not a desirable feature, we know from the problem of collinearity that the model with all potential independent variables included may not even be viable, let alone the best amongst alternatives. The examples of model selection we have encountered involved only a small number of potential independent variables, allowing the analysis of the alternatives to be conducted in an ad-hoc manner.

However, in practical situations we often have to choose the independent variables from a very large set of potential variables.

In these situations, an ad hoc approach is infeasible because of the sheer number of alternative models to consider.

We need a **system**, a set of rules and instructions that when followed lead to the selection of the **best** model from possible alternatives.

Even better would be a system that can be **automated**.

Consider a task involving the identification of the best-performing model from those that can be constructed with combinations of independent variables selected from a set of size m.

Let \boldsymbol{x} be the *m*-dimensional vector of possible independent variables.

The best model could be one of many, ranging from the null model

$$\hat{Y}|\boldsymbol{x} = \hat{\beta}_0$$

to the **maximal model**

$$\hat{Y}|\mathbf{x} = \hat{\beta}_0 + \sum_{j=1}^m \hat{\beta}_j x_j$$

and every simple and multiple regression model in between.

The problem is that there are 2^m possible combinations of independent variables and therefore 2^m possible regression models to test.

The number of possible models grows **exponentially** as the dimension m increases.

This is an example of the **curse of dimensionality** that renders some numerical tools, proven in low dimensions, infeasible in high dimensions.

Each of the model selection systems we consider has a method of navigating, or **iterating**, through the possible models and for comparing one against another.

The process of model selection relies to a large extent on theory already developed, so we illustrate the main ideas with an example

Consider the problem of finding the best performing regression model for predicting school-average student test performance.

The response (or dependent variable) is

mean student test score (Y or score)

and potential predictors (independent variables)

- staff salaries per pupil (*x*₁ or *sal*)
- % white-collar fathers (x₂ or dad)
- socioeconomic status (*x*₃ or *ses*)
- teachers' mean score (*x*₄ or *teach*)
- mothers' mean education level (x_5 or mumed).

Data is available in verbal.csv on Canvas.

5. School performance example - setup

After collecting the sample data $(X_{i,1}, \ldots, X_{i,5}, Y_i)$, we look to characterise the relationships between the variables, hoping to spot linear relationships as we seek to fit a plane.



5. School performance example - setup

From the scatter plot we spot seemingly strong positive relationships between the sample data dependent variable Y_i (*score_i*) and sample data independent variables $X_{i,2}$ (*dad_i*), $X_{1,3}$ (*ses_i*), and $X_{i,5}$ (*mumed*).

This visual analysis is confirmed by correlation analysis of the sample data

$$\operatorname{corr}(Y_i, X_{i,1}) = \operatorname{corr}(score_i, sal_i) = 0.192$$

$$\operatorname{corr}(Y_i, X_{i,2}) = \operatorname{corr}(score_i, dad_i) = 0.753$$

$$\operatorname{corr}(Y_i, X_{i,3}) = \operatorname{corr}(score_i, ses_i) = 0.927$$

$$\operatorname{corr}(Y_i, X_{i,4}) = \operatorname{corr}(score_i, teach_i) = 0.334$$

$$\operatorname{corr}(Y_i, X_{i,5}) = \operatorname{corr}(score_i, mumed_i) = 0.732.$$

Seeing potential for **collinearity** between x_2 (*dad*) and x_3 (*ses*), we look at their sample correlation

$$\operatorname{corr}(X_{i,2}, X_{i,3}) = \operatorname{corr}(dad_i, ses_i) = 0.827$$

which although high, may not be high enough to trigger a *VIF* (variance inflation factor) warning.

In all there are $2^5 = 32$ possible models, ranging from the null model to the maximal model.

We look at four possible selection methods:

- selection from all possible models
- forward selection
- backward elimination
- stepwise regression.

The obvious way to ensure the best performing model from a set is identified is to test them all. Using the strict interpretation of the **selection from all** method, the best performing model is the one with the best **ranking statistic**.

The ultimate barrier to this method is the curse of dimensionality mentioned earlier, but with only 32 models to choose from this is not an issue here.

Even when feasible, this method still produces many models for analysis with no prescriptive method to do so.

If not applied strictly, many statistical properties of each model must be compared which, given the often conflicting information they provide, can lead to selection being performed in an ad hoc or subjective manner.

In this example we use $R_{\rm adj}^2$ and the p-value from the F-test as ranking statistics.

Data from all models (but null) is summarised below.

```
> library('olsrr')
> mod1 <- lm(score ~ ., data = verbal)</pre>
> ols step all possible(mod1)
Index N
                   Predictors
                               R-Square Adj. R-Square Mallow's Cp
3
      1 1
                            ses 0.85962771 0.851829254 5.002821
2
  2 1
                            dad 0.56761278 0.543591266 48.694760
5
   31
                          mumed 0.53571268 0.509918939 53.467726
4
   4 1
                         teach 0.11132200 0.061950995 116.966026
1
   51
                            sal 0.03697606 -0.016525270 128.089834
13
      6 2
                      ses teach 0.88734850 0.874095385 2.855174
14 7 2
                      ses mumed 0.86180687 0.845548856 6.676771
10 8 2
                        dad ses 0.86020753
                                           0.843761359 6.916067
7
    92
                        sal ses 0.86007615 0.843614519 6.935725
   10 2
                      dad teach 0.65497310 0.614381699
                                                        37,623710
11
15
    11 2
                     teach mumed 0.59599405
                                           0.548463939
                                                        46,448290
12
    12 2
                      dad mumed 0.57561917 0.525692012
                                                        49,496826
6 13.2
                        sal dad 0.57083444
                                           0.520344375
                                                        50.212728
9
   14 2
                      sal mumed 0.53820829 0.483879856
                                                        55.094327
8
    15 2
                     sal teach 0.11213031
                                           0.007675058
                                                        118,845084
```

19	16 3	sal ses teach 0.90071101	0.882094325	2.855845
25	17 3	ses teach mumed 0.88884972	0.868009037	4.630559
22	18 3	dad ses teach 0.88738875	0.866274139	4.849152
20	19 3	sal ses mumed 0.86222800	0.836395753	8.613760
23	20 3	dad ses mumed 0.86216960	0.836326405	8.622498
16	21 3	sal dad ses 0.86067265	0.834548777	8.846475
17	22 3	sal dad teach 0.66622665	0.603644148	37.939928
24	23 3	dad teach mumed 0.65590099	0.591382431	39.484877
21	24 3	sal teach mumed 0.60260264	0.528090632	47.459498
18	25 3	sal dad mumed 0.57813337	0.499033377	51.120646
29	26 4	sal ses teach mumed 0.90185831	0.875687195	4.684183
26	27 4	sal dad ses teach 0.90096690	0.874558069	4.817558
30	28 4	dad ses teach mumed 0.89224918	0.863515622	6.121924
27	29 4	sal dad ses mumed 0.86255877	0.825907778	10.564269
28	30 4	sal dad teach mumed 0.66696870	0.578160357	39.828901
31	31 5 sal	l dad ses teach mumed 0.90643105	0.873013562	6.000000

Note:

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- You can request the coefficients in each model by using ols_step_all_possible_betas(mod1)
- We can also use regsubsets function in leaps package which gives the best subset of each size according to some selection criteria.

Let's see how far we get selecting the best performing model as the regression with the highest R_{adj}^2 with *p* value less than 0.05 (regression is significant).

1st Choice: $R_{adj}^2 = 0.882$

Coefficients:

Estimate Std. Error t value Pr(>|t|) (Intercept) 12.11951 9.03643 1.341 0.199 sal -1.73581 1.18290 -1.467 0.162 ses 0.55321 0.04907 11.273 5.06e-09 *** teach 1.03582 0.40479 2.559 0.021 * ---Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '

Residual standard error: 1.997 on 16 degrees of freedom Multiple R-squared: 0.9007,Adjusted R-squared: 0.8821 F-statistic: 48.38 on 3 and 16 DF, p-value: 3.011e-08

If not we have to decide what to do with x_1 (*sal*), which is insignificant (p = 0.162). Should it be removed or does it contribute to the regression assumptions being satisfied?

2nd Choice: $R_{adi}^2 = 0.876$

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	15.52853	12.34376	1.258	0.2276	
sal	-1.71421	1.21571	-1.410	0.1789	
ses	0.58246	0.08612	6.763	6.38e-06	***
teach	1.02494	0.41645	2.461	0.0265	*
mumed	-0.52545	1.25481	-0.419	0.6813	
Signif. cod	es: 0'**	*' 0.001'*	**' 0.01	'*' 0.05	'.' 0.1'

Residual standard error: 2.051 on 15 degrees of freedom Multiple R-squared: 0.9019,Adjusted R-squared: 0.8757 F-statistic: 34.46 on 4 and 15 DF, p-value: 2.133e-07

This model is next in the ranking, but now we have two variables, x_1 (*sal*) and x_5 (*mumed*), that are insignificant. This is hardly an improvement over the previous model, which had only one insignificant variable and a higher R_{adi}^2 .

Note that this model is formed by adding a variable to the 1st Choice model.

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3rd Choice: R_{adj}^2 = 0.875
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Coefficients:

	Estimate	Std. Error †	t value	Pr(> t)	
(Intercept)	11.535220	9.781866	1.179	0.257	
sal	-1.755821	1.224347	-1.434	0.172	
ses	0.538487	0.090308	5.963	2.6e-05	***
teach	1.052506	0.426037	2.470	0.026	*
dad	0.006525	0.033145	0.197	0.847	
Signif. code	es: 0'***	, 0.001 (**	, 0.01 ʻ	*' 0.05'	.' 0.1 ' '

Residual standard error: 2.06 on 15 degrees of freedom Multiple R-squared: 0.901,Adjusted R-squared: 0.8746 F-statistic: 34.12 on 4 and 15 DF, p-value: 2.281e-07

This model is third in the ranking, but again we have two variables, this time x_1 (sal) and x_2 (dad), that are insignificant. This model is inferior to the 1st Choice.

Note that this model is formed by adding a variable to the 1st Choice model.

4th Choice: $R_{adj}^2 = 0.874$

Coefficients:

Estimate Std. Error t value Pr(>|t|) (Intercept) 14.58268 9.17541 1.589 0.1304 ses 0.54156 0.05004 10.822 4.81e-09 *** teach 0.74989 0.36664 2.045 0.0566 . ---Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '

Residual standard error: 2.064 on 17 degrees of freedom Multiple R-squared: 0.8873,Adjusted R-squared: 0.8741 F-statistic: 66.95 on 2 and 17 DF, p-value: 8.705e-09

Now looking at the fourth on the list, once more there is an insignificant variable, x_4 (*teach*), albeit with a p-value not too far from the 0.05 significance level. (p = 0.057). But if we are going break rules we might as well stay with the 1st Choice model.

Note that this model is formed by removing a variable from the 1st Choice model.

5th Choice: $R_{adj}^2 = 0.852$

Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 33.32280 0.52800 63.11 < 2e-16 *** ses 0.56033 0.05337 10.50 4.2e-09 *** ---Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '

Residual standard error: 2.239 on 18 degrees of freedom Multiple R-squared: 0.8596,Adjusted R-squared: 0.8518 F-statistic: 110.2 on 1 and 18 DF, p-value: 4.199e-09

It has taken until fifth in the ranking before finding a model with no insignificant variables, and this one is a simple linear regression.

Note that this model is formed by removing a variable from the 4th Choice model.

What can be learned from attempting to use a measure like R_{adj}^2 to rank the performance of all models in this example?

- **E**asy to be in a situation requiring some statistics to be either ignored or treated inconsistently.
- 2 Moving between models in the ranking involved adding or removing independent variables.

The second point hints at a different approach to model selection, one based on choosing a particular model as a starting point (often the null or maximal model) and adding or deleting independent variables until some condition is met.

The final stop in such an iterative process is the model selected as best performing.

One such approach could be to start with the significant simple regression model with the highest R_{adj}^2 . At each iteration the variable associated with the largest increase in R_{adj}^2 (keeping the regression significant) is added to the model. The process stops the first time R_{adi}^2 falls.

The models considered under this scheme are summarised below.

				R_sq	Ad_R_Sq	p-value (F)
ses (0.000)				0.860	0.852	0.000
dad				0.568	0.544	0.000
mumed				0.536	0.510	0.000
teach				0.111	0.062	0.151
sal				0.037		0.417
ses (0.000)	teach (0.057)			0.887	0.874	0.000
ses	mumed			0.862	0.846	0.000
ses	sal			0.860	0.844	0.000
ses	dad			0.860	0.844	0.000
ses (0.000)	teach (0.021)	sal (0.162)		0.901	0.882	0.000
ses	teach	mumed		0.889	0.868	0.000
ses	teach	dad		0.887	0.866	0.000
ses (0.000)	teach (0.026)	sal (0.179)	mumed (0.681)	0.902	0.876	0.000
ses	teach	sal	dad	0.901	0.875	0.000

Note: t-stats in brackets

This method involves selection from a subset of all regressions.

The principle argument used against the selection from all method becomes weaker as processing power increases. However, in activities where processing power is already fully deployed (e.g. quantitative trading), this argument still stands.

Rather than R_{adj}^2 , the selection from all method can use Mallows C_p statistic, Akaike Information Criterion (AIC) or Bayseian Information Criterion (BIC) as the ranking statistic.

Don't use the selection from all method with R^2 as the ranking statistic. This results in the maximal model being selected (recall that R^2 increases with m). The first method is **forward selection**, where a provisionally optimal model \hat{Y}^* is augmented with the most significant additional predictor at each iteration.

It is defined by the following sequential steps:

- **1** set \hat{Y}^* as null model (no predictors)
- **2** construct all possible models by adding one predictor to \hat{Y}^*
 - if all models are insignificant (F-test) GO TO 3
 - if not set \hat{Y}^* as model with most significant new predictor (T-test) and **REPEAT** 2

3 the optimal model is current iteration of \hat{Y}^* . **STOP**.

In R, we can use the command ols_step_forward_p(mod1, penter = 0.5, details = T) to perform forward selection. Using this method, the model selected has independent variable x_3 (ses) with $R_{adj}^2 = 0.852$, the 5th Choice using selection from all.

Final Model Output

model	Beta	Std. Error	Std. Beta		Sig	lower	upper
(Intercept)	33.323	0.528	0.927	63.112	0.000	32.214	34.432
ses	0.560	0.053		10.499	0.000	0.448	0.672

Parameter Estimates

Another method is **backward selection**, where a provisionally optimal model \hat{Y}^* is contracted by removing the least significant predictor at each iteration.

It is defined by the following sequential steps:

1 set \hat{Y}^* as maximal model (all predictors)

2 identify least significant predictor (T-test) in \hat{Y}^*

- if all are significant GO TO 3
- if not construct model by removing identified predictor, set \hat{Y}^* as this model and REPEAT [2]

3 the optimal model is current iteration of \hat{Y}^* . **STOP**.

In R, we can use the command ols_step_backward_p(mod1, prem = 0.1, details=T) to perform forward selection. Using this method, the model selected has independent variables x_3 (ses) and x_4 (teach) with $R_{adj}^2 = 0.874$, the 4th Choice using selection from all.

Final Model Output									
Parameter Estimates									
model	Beta	Std. Error	Std. Beta	t	Sig	lower	upper		
(Intercept) ses teach	14.583 0.542 0.750	9.175 0.050 0.367	0.896 0.169	1.589 10.822 2.045	0.130 0.000 0.057	-4.776 0.436 -0.024	33.941 0.647 1.523		

Elimination Summary							
Step	Variable Removed	R-Square	Adj. R-Square	C(p)	AIC	RMSE	
1 2 3	dad mumed sal	0.9019 0.9007 0.8873	0.8757 0.8821 0.8741	4.6842 2.8558 2.8552	91.7366 89.9690 90.4943	2.0510 1.9974 2.0641	

Note that we remove variables if p > 0.1 (prem=0.1). Setting this to the more conservative p > 0.050001 (removal threshold must be set greater than significance level) results in the selection of the model with independent variable x_3 (ses) and $R_{adi}^2 = 0.852$, the 5th Choice using selection from all and the same model using forward selection.

Elimination Summary							
Step	Variable Removed	R-Square	Adj. R-Square	C(p)	AIC	RMSE	
1	dad	0.9019	0.8757	4.6842	91.7366	2.0510	
2	mumed	0.9007	0.8821	2.8558	89.9690	1.9974	
3	sal	0.8873	0.8741	2.8552	90.4943	2.0641	
4	teach	0.8596	0.8518	5.0028	92.8943	2.2392	

Elimination Cummons

The final iterative method is stepwise regression.

It is defined by the following sequential steps:

- **1** set \hat{Y}^* as chosen initial model (often null or maximal)
- **2** construct all possible models by adding one predictor to \hat{Y}^*
 - if all models are insignificant (F-test) GO TO 4
 - if not set \hat{Y}^* as model with most significant new predictor (T-test) and **GO TO** 3

 ${\tt 3}$ identify least significant predictor (T-test) in \hat{Y}^*

- if all are significant GO TO 4
- if not construct model by removing identified predictor, set \hat{Y}^* as this model and GO TO $\space{2}$

4 the optimal model is current iteration of \hat{Y}^* . **STOP**.

In R, we can use the command ols_step_both_p(mod1, pent = 0.05, prem = 0.1, details = T)) to perform stepwise selection.

Using this method, the model selected has independent variable x_3 (ses) with $R_{adj}^2 = 0.852$, the 5th Choice using selection from all, also the model selected using forward selection and also the model using backward selection (if default exit threshold is lowered from 0.1).

Parameter Estimates							
model	Beta	Std. Error	Std. Beta	t	Sig	lower	upper
(Intercept) ses	33.323 0.560	0.528 0.053	0.927	63.112 10.499	0.000 0.000	32.214 0.448	34.432 0.672

Draper, N. R. and Smith, H. (1998). *Applied regression analysis*. Wiley-Interscience, Somerset, US.